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Does Causality Technique Matter to Savings-Growth Nexus in Malaysia?

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ABSTRACT

The intention of this study was to investigate whether the causal inference between savings and economic growth in Malaysia is sensitive to the particular causality tests employed to ascertain the causal relationship. This study covered quarterly data from 1991:Q1 to 2006:Q3. The results suggested that the causal relationship between savings and economic growth in Malaysia is not sensitive to the particular causality test used. Thus, causality test plays no role in explaining the inconsistency causality result of savings and economic growth. Ultimately, causality test does not matter to savings-growth nexus for Malaysia.

Keywords: Causality; parametric; nonparametric; savings; growth.

JEL Classification Code: C14; C22; E21; O16.

ABSTRAK

Tujuan kajian ini dijalankan adalah untuk mengkaji sama ada hubungan sebab-penyebab antara tabungan dan pertumbuhan ekonomi di Malaysia adalah sensitif terhadap pendekatan yang digunakan untuk menentukan arah sebab-penyebab. Kajian ini menggunakan data sukuan dari tahun 1991:1 hingga 2006:3. Keputusan empirikal kajian ini mendapati bahawa hubungan sebab-penyebab antara tabungan dan pertumbuhan ekonomi di Malaysia tidak sensitif terhadap pendekatan sebab-penyebab yang digunakan. Justeru, pendekatan sebab-penyebab bukan penentu kepada ketidakserasian hubungan sebab-penyebab antara tabungan dan pertumbuhan ekonomi di Malaysia.

Kata kunci: Sebab-penyebab; parametrik; bukan parametrik; tabungan; pertumbuhan ekonomi.

Klasifikasi Kod JEL: C14; C22; E21; O16.

INTRODUCTION

By the beginning of the 21st century, Malaysia had become one of the fastest growing economies in the Southeast Asian region.¹ Various theoretical explanations have been provided by many researchers to understand the reason behind this impressive growth. One of the interesting areas that has not been provided with consensus evidence is the causal link between savings and economic

growth. This causal relationship is also known as the savings-growth nexus. Many researchers have empirically investigated the causal link between savings and economic growth in Malaysia through various model specifications (Gruben & McLeod, 1998; Agrawal, 2001; Baharumshah, Thanoon, & Rashid, 2003). Unfortunately, previous empirical tests failed to produce clear evidence of the causal link. In view of literature, some empirical studies claimed that economic growth causes savings to

change, (Sinha & Sinha, 1998; Carroll, Overland, & Weil, 2000; Rodrik, 2000; Baharumshah, et al., 2003), however others defended the view that savings induce economic growth through its impact on capital formation (Lewis, 1955; Levine & Renelt, 1992; Mankiw, Romer, & Weil, 1992; Agrawal, 2001; Alguacil, Cuadros, & Orts, 2004; Tang, 2008). This controversial finding has sparked the interest of this study to further investigate the issue of savings-growth nexus in Malaysia from a different perspective.

Chowdury (1987) argued that in most cases the non-consensus causality result was probably attributed to the different causality tests employed. Thornton and Batten (1985) and Xu (1996) added that the arbitrariness in the choice of lag order for causality test may also yield different causality results. Moreover, Lee, Lin, and Wu (2002) conducted a Monte Carlo experiment to examine the performance of Granger causality test in detecting the true causal relationship. They pointed out that relying on one causality test may not be enough to identify the true causal relationship. As far as Malaysia is concerned, empirical studies for the savings-growth nexus are based on a single causality test to verify the causality direction for formulation and implementation of appropriate macroeconomics policies. For this reason, the policy suggestion based on one causality test is not without question. Therefore, the goal of this study is to re-investigate the causality direction between savings and economic growth in Malaysia through parametric and nonparametric causality testing procedures.² By testing the causality direction with various methods, we were able to assess whether savings-growth nexus in Malaysia is sensitive to the particular causality test used in determining the causal relationship. To the best of our knowledge, none of the existing empirical studies has tested the role of causality techniques on the issue of savings-growth nexus. In addition to that, the findings of this study that are based on various causality tests provides a more reliable result to the policy maker, whether freeing up domestic resources such as savings is helpful or harmful to the Malaysian economy. This econometric exercise is parallel to the suggestion noted in Lee et al. (2002). If the causality evidence

suggests that savings Granger causes economic growth, this reflects that the financial system in Malaysia successfully translates savings into capital formation which fosters economic growth. Otherwise, if economic growth is not the result of savings, policy initiatives that encourage savings could be detrimental to the economy.

The remainder of this article is set out as follows. Section will briefly discuss the data, model, and stationarity test. Section will shortly present the causality testing procedures use in this study. Section reports the empirical results and Section concludes.

DATA, MODEL, AND TATIONARITY TEST

Data and Model

This study uses quarterly data of gross domestic saving (GDS), gross domestic product (GDP), and consumer price index (CPI, 2000 = 100) from 1991:Q1 to 2006:Q3 in Malaysia extracted from International Monetary Fund (IMF) publication, *International Financial Statistics* (IFS) and Bank Negara Malaysia publication, *Monthly Statistical Bulletin*. The estimated variables are deflated by CPI to obtain the real terms. Quarterly data is used in this article in order to yield more power on statistical test and to avoid the size distortion problem (Zhou, 2001).

In this article, we use the bivariate savings function that relates to economic growth. The savings function can be written as:

$$\ln S_t = \alpha + \beta \ln G_t + \varepsilon_t \quad (1)$$

where $\ln S_t$ represents the natural log of real gross domestic saving (*RGDS*) and $\ln G_t$ is the natural log of real gross domestic product (*RGDP*); ε_t is an error term assumed to be white noise and complies to the classical assumption.

Stationarity Test

It is well known that the standard Augmented Dickey-Fuller (ADF) and Phillips-Perron (PP) tests have been criticised for having low power

in distinguishing between unit root and a near unit root stationary process (Campbell & Perron, 1991; DeJong, Nankervis, Savin, & Whiteman, 1992). The low power of ADF and PP unit root tests has prompted the present study to use the null stationary test proposed by Kwiatkowski, Phillips, Schmidt and Shin (1992) – KPSS. The KPSS semi-parametric approach is to test the null hypothesis of level (η_μ) or trend (η_τ) stationarity against the alternative hypothesis of a unit root. The following expresses the KPSS test statistics equation,

$$LM = \eta_\mu(\eta_\tau) = \frac{1}{s^2(k)T^2} \sum_{t=1}^T S_t^2 \quad (2)$$

where $S_t = \sum_{i=1}^t u_i$, u_t are residuals from the OLS regression of y_t on the exogenous variables x_t . $s^2(k)$ is the nonparametric estimate of the long run variance of u_t and k represents for the lag truncation parameter.

CAUSALITY TESTING PROCEDURES

In this section, we briefly discuss the selected five causality testing procedures.

Parametric Causality Techniques

(i) Granger Test

Since the standard Granger (1969) causality testing procedure has been widely used in earlier empirical studies, only a brief discussion is offered here. Causality in the Granger sense asserts that savings causes economic growth if the past values of savings can be used to forecast economic growth more accurately than just the past values of economic growth. Therefore, the Granger causality testing equation can be expressed as follows.

$$\Delta \ln G_t = \alpha + \sum_{i=1}^p \delta_i \Delta \ln G_{t-i} + \sum_{j=1}^q \lambda_j \Delta \ln S_{t-j} + \varepsilon_{1t} \quad (3)$$

$$\Delta \ln S_t = \alpha + \sum_{i=1}^p \lambda_i \Delta \ln S_{t-i} + \sum_{j=1}^q \delta_j \Delta \ln G_{t-j} + \varepsilon_{2t} \quad (4)$$

where Δ is the first difference operator; p and q are the maximum lag order, and the residuals ($\varepsilon_{1t}, \varepsilon_{2t}$) are assumed to be white noise and spherically distributed. The Granger causality test is implemented by computing the F -statistic on the lagged variables. From equation (3), $\lambda_j \neq 0 \forall_j$ implies that there is causality running from savings to economic growth; whereas in equation (4), economic growth Granger causes savings, if $\delta_j \neq 0 \forall_j$ holds.

(ii) Modified Sims Test

As an alternative to Granger causality test, Sims (1972) developed a causality test based on the general concept of causality that the “future cannot cause the present”. This test takes into account the past, current, and future values of regressor to examine the causal link between two variables. Unfortunately, the residuals in the Sims causality test are generally serially correlated. Therefore, eweke, Meese and Dent (1983) proposed to incorporate the lagged dependent variable(s) into the testing equation to overcome the serial correlation problem. The modified Sims test can be performed by estimating the following equations,

$$\Delta \ln G_t = \alpha + \sum_{i=1}^p \lambda_i \Delta \ln G_{t-i} + \sum_{j=0}^q \delta_j \Delta \ln S_{t-j} + \sum_{l=1}^r \phi_l \Delta \ln S_{t+l} + \varepsilon_{1t} \quad (5)$$

$$\Delta \ln S_t = \alpha + \sum_{i=1}^p \lambda_i \Delta \ln S_{t-i} + \sum_{j=0}^q \delta_j \Delta \ln G_{t-j} + \sum_{l=1}^r \phi_l \Delta \ln G_{t+l} + \varepsilon_{2t} \quad (6)$$

where Δ is the first difference operator; p and q are the maximum lag order and r represent the maximum lead order. The residuals ($\varepsilon_{1t}, \varepsilon_{2t}$) are assumed to be white noise and spherically distributed. The different between Sims and modified Sims causality tests is the inclusion

of lagged dependent variable(s) to remove the serial correlation problem. However, this method is slightly different from the standard Granger causality test, particularly in the testing procedure. In order to examine whether economic growth causes savings, we estimated equation (5) and then employed the F -statistic on the null hypothesis of non-causal relation ($H_0: \varphi_1 = \varphi_2 = \dots = \varphi_i = 0$).³ If the null hypothesis is rejected, this implies that economic growth causes savings to change. Otherwise, economic growth does not cause savings to change. Inversely, we estimated equation (6) to ascertain the causality from savings to economic growth. If the null hypothesis of non-causal relation ($H_0: \theta_1 = \theta_2 = \dots = \theta_i = 0$) is rejected, this asserts that savings causes economic growth to change, otherwise the null hypothesis of savings does not cause economic growth is true.

(iii) Hsiao Test

The literature stated that the Granger causality tests are sensitive to the lag order in the autoregressive process. An inadequate choice of lag length would lead to inconsistent model estimates and the causality inference would likely be misleading. Responding to this, Hsiao (1981) proposed a causality test by combining the Granger concept of causality and information criteria to avoid imposing false and spurious restrictions on the model.⁴ This causality testing approach is based on the idea that the inclusion of causal variable(s) into the estimated model will reduce the size of variation and hence increase the ability to forecast. Hsiao's version of Granger causality test is a two-step procedure to determine the optimal lag lengths and the direction of causality at the same time. In order to ascertain the null hypothesis of savings does not Granger cause economic growth to change, we consider the following model equations,

$$\Delta \ln G_t = \alpha + \sum_{i=1}^p \delta_i \Delta \ln G_{t-i} + \varepsilon_{1t} \quad (7)$$

$$\Delta \ln G_t = \alpha + \sum_{i=1}^p \lambda_i \Delta \ln G_{t-i} + \sum_{j=1}^q \delta_j \Delta \ln S_{t-j} + \varepsilon_{2t} \quad (8)$$

where Δ is the first difference operator. The p and q represents the maximum lag order. The error terms $(\varepsilon_{1t}, \varepsilon_{2t})$ are assumed to be normally distributed and serially uncorrelated. In Hsiao's test, the first step is computing the autoregressive (AR) process model as given by equation (7). Then the Akaike's Information Criterion (AIC) is computed with the order of lags i varying from 1 to p . The lag p with the smallest AIC values is chosen and denotes the corresponding AIC as $AIC(p,0)$.

In the second step, $\Delta \ln G_t$ is treated as controlled variables with p lags, then lags of $\Delta \ln S_t$ are added sequentially into equation (8). After that, the AIC is computed with the order of lags j varying from 1 to q . The lag q with the smallest AIC values is chosen and denote the corresponding AIC as $AIC(p,q)$. According to Hsiao's (1981) interpretation; if $AIC(p,0) > AIC(p,q)$ then $\Delta \ln S_t$ Granger causes $\Delta \ln G_t$. On the other hand, if $AIC(p,0) < AIC(p,q)$ then $\Delta \ln S_t$ does not Granger cause $\Delta \ln G_t$.

(iv) Modified Wald (MWALD) Test

He and Maekawa (2001) argued that F -statistics for the Granger causality test often leads to spurious causality inference when one or both of the estimated series are non-stationary. Furthermore, the low power unit root test will always have a degree of uncertainty with respect to the order of integration. Due to this problem, Toda and Yamamoto (1995) proposed a simple procedure (modified Wald – MWALD test) which involves the estimation of an augmented vector autoregression (VAR) model at level irrespective of the order of integration. Thus, pre-testing of unit root is not required in this causality test. In short, the test will conduct with variables at level by adding extra lag (d_{\max})

into the VAR system to ensure that the asymptotic critical values can be applied when the test is conducted between the integrated variables. The d_{\max} refers to the suspected maximum order of integration. Thus, the MWALD testing equation can be expressed as follows,

$$\ln S_t = \alpha + \sum_{i=1}^p \lambda_i \ln S_{t-i} + \sum_{j=p+1}^{d_{\max}} \pi_j \ln S_{t-j} + \sum_{i=1}^q \delta_i \ln G_{t-i} + \sum_{j=q+1}^{d_{\max}} \theta_j \ln G_{t-j} + \varepsilon_{1t} \quad (9)$$

$$\ln G_t = \alpha + \sum_{i=1}^p \delta_i \ln G_{t-i} + \sum_{j=p+1}^{d_{\max}} \theta_j \ln G_{t-j} + \sum_{i=1}^q \lambda_i \ln S_{t-i} + \sum_{j=q+1}^{d_{\max}} \pi_j \ln S_{t-j} + \varepsilon_{2t} \quad (10)$$

In order to employ the MWALD test, we have to pre-specify the maximum order of integration (d_{\max}) for the series in the VAR system. Regarding the extra lag (d_{\max}) variables, Dolado and Lütkepohl (1996) suggested using $d_{\max} = 1$ due to its best performance in their Monte Carlo study. In this respect, we used $d_{\max} = 1$ in our study.⁵ Similarly, the F -statistic is employed to examine the existence of causal relationship. From equation (9), $\delta_i \neq 0 \forall_i$ asserts that there is causality running from economic growth to savings; whereas in equation (10), savings causes economic growth, if $\lambda_i \neq 0 \forall_i$ holds.

Nonparametric Causality Technique

Toward this end, our study investigated the savings-growth nexus in Malaysia using nonparametric causality methodology because of its superior statistical properties to the parametric approach. Holmes and Hutton (1988, 1990a) argued that the parametric causality tests (i.e. Granger, Sims, modified Sims, Hsiao, and MWALD) may be problematic compared to nonparametric procedures. This is because the parametric causality approaches are based on the maintained hypotheses of correct functional form, homoskedasticity, the residuals are free from the serial correlation problem, and data are

normally distributed. Violation of these classical assumptions will affect the causality inferences (Gordon & Sakyi-Bekoe, 1993; Ansari, Gordon, & Akuamoah, 1997; Tang & Lean, 2008). Thus, Holmes and Hutton (1988, 1990a) proposed an alternative Granger causality testing procedure based on the rank ordering (R) of each variable. That is, they suggested ranking each variable and using the rank value of each observation to test for causality. If the maintained hypothesis is violated this nonparametric approach is robust over the alternative parametric Granger causality test. Holmes and Hutton (1990b) added that the nonparametric approach multiple rank F -test is approximately twice as powerful as the parametric approach, when the sample size is small and linear. The multiple rank F -test equation can be written as follows:

$$R(\Delta \ln S_t) = \alpha + \sum_{i=1}^p \lambda_i R(\Delta \ln S_{t-i}) + \sum_{j=1}^q \delta_j R(\Delta \ln G_{t-j}) + \varepsilon_{1t} \quad (11)$$

$$R(\Delta \ln G_t) = \alpha + \sum_{i=1}^p \delta_i R(\Delta \ln G_{t-i}) + \sum_{j=1}^q \lambda_j R(\Delta \ln S_{t-j}) + \varepsilon_{2t} \quad (12)$$

where $R(\cdot)$ represents a rank order transformation; p and q refer to the maximum lag length incorporated into the testing equation. The residuals ε_{1t} and ε_{2t} are assumed to be serially uncorrelated and white noise. Then, the F -statistic is employed to examine the possibility of presence of a causal relationship between savings and economic growth. From equation (11), $\delta_j \neq 0 \forall_j$ implies that there is causality from economic growth to savings; whereas in equation (12), savings causes economic growth, if $\lambda_j \neq 0 \forall_j$.

Table 1 and Table 2 summarise the five causality testing equation and hypothesis respectively.

Table 1: Five Types of Causality Testing Model

Types of Causality Tests	Causality Testing Model	
	Economic Growth (G_t) causes Savings (S_t)	Savings (S_t) causes Economic Growth (G_t)
Granger	$\Delta \ln S_t = \alpha + \sum_{i=1}^p \lambda_i \Delta \ln S_{t-i} + \sum_{j=1}^q \delta_j \Delta \ln G_{t-j} + \varepsilon_t$	$\Delta \ln G_t = \alpha + \sum_{i=1}^p \delta_i \Delta \ln G_{t-i} + \sum_{j=1}^q \lambda_j \Delta \ln S_{t-j} + \varepsilon_t$
Modified Sims	$\Delta \ln G_t = \alpha + \sum_{i=1}^p \lambda_i \Delta \ln G_{t-i} + \sum_{j=0}^q \delta_j \Delta \ln S_{t-j} + \sum_{l=1}^r \phi_l \Delta \ln S_{t+l} + \varepsilon_t$	$\Delta \ln S_t = \alpha + \sum_{i=1}^p \lambda_i \Delta \ln S_{t-i} + \sum_{j=0}^q \delta_j \Delta \ln G_{t-j} + \sum_{l=1}^r \phi_l \Delta \ln G_{t+l} + \varepsilon_t$
Hsiao	$\Delta \ln S_t = \alpha + \sum_{i=1}^p \lambda_i \Delta \ln S_{t-i} + \sum_{j=1}^q \delta_j \Delta \ln G_{t-j} + \varepsilon_t$	$\Delta \ln G_t = \alpha + \sum_{i=1}^p \delta_i \Delta \ln G_{t-i} + \sum_{j=1}^q \lambda_j \Delta \ln S_{t-j} + \varepsilon_t$
Modified Wald	$\ln S_t = \alpha + \sum_{i=1}^p \lambda_i \ln S_{t-i} + \sum_{j=p+1}^{d_{\max}} \pi_j \ln S_{t-j} + \sum_{i=1}^q \delta_i \ln G_{t-i} + \sum_{j=q+1}^{d_{\max}} \theta_j \ln G_{t-j} + \varepsilon_t$	$\ln G_t = \alpha + \sum_{i=1}^p \delta_i \ln G_{t-i} + \sum_{j=q+1}^{d_{\max}} \theta_j \ln G_{t-j} + \sum_{i=1}^q \lambda_i \ln S_{t-i} + \sum_{j=p+1}^{d_{\max}} \pi_j \ln S_{t-j} + \varepsilon_t$
Multiple Rank F-test	$R(\Delta \ln S_t) = \alpha + \sum_{i=1}^p \lambda_i R(\Delta \ln S_{t-i}) + \sum_{j=1}^q \delta_j R(\Delta \ln G_{t-j}) + \varepsilon_t$	$R(\Delta \ln G_t) = \alpha + \sum_{i=1}^p \delta_i R(\Delta \ln G_{t-i}) + \sum_{j=1}^q \lambda_j R(\Delta \ln S_{t-j}) + \varepsilon_t$

Table 2: Hypothesis Testing and AIC Criteria for Causality in Five Models

Types of Causality Tests	Unidirectional		Bilateral	No Causality
	$G_t \rightarrow S_t$	$S_t \rightarrow G_t$	$S_t \leftrightarrow G_t$	$S_t \nleftrightarrow G_t$
Granger	$H_0 : \sum_{j=1}^q \delta_j \neq 0$	$H_0 : \sum_{j=1}^q \lambda_j \neq 0$	$H_0 : \sum_{j=1}^q \delta_j \neq 0$ $H_0 : \sum_{j=1}^q \lambda_j \neq 0$	$H_0 : \sum_{j=1}^q \delta_j = 0$ $H_0 : \sum_{j=1}^q \lambda_j = 0$
Modified Sims	$H_0 : \sum_{l=1}^r \phi_l \neq 0$	$H_0 : \sum_{l=1}^r \phi_l \neq 0$	$H_0 : \sum_{l=1}^r \phi_l \neq 0$ $H_0 : \sum_{l=1}^r \phi_l \neq 0$	$H_0 : \sum_{l=1}^r \phi_l = 0$ $H_0 : \sum_{l=1}^r \phi_l = 0$
Hsiao	$AIC(\Delta \ln S)$ $> AIC(\Delta \ln S, \Delta \ln G)$	$AIC(\Delta \ln G)$ $> AIC(\Delta \ln G, \Delta \ln S)$	$AIC(\Delta \ln S)$ $> AIC(\Delta \ln S, \Delta \ln G);$ $AIC(\Delta \ln G)$ $> AIC(\Delta \ln G, \Delta \ln S)$	$AIC(\Delta \ln S)$ $= AIC(\Delta \ln S, \Delta \ln G);$ $AIC(\Delta \ln G)$ $= AIC(\Delta \ln G, \Delta \ln S)$
Modified Wald	$H_0 : \sum_{i=1}^q \delta_i \neq 0$	$H_0 : \sum_{i=1}^q \lambda_i \neq 0$	$H_0 : \sum_{i=1}^q \delta_i \neq 0$ $H_0 : \sum_{i=1}^q \lambda_i \neq 0$	$H_0 : \sum_{i=1}^q \delta_i = 0$ $H_0 : \sum_{i=1}^q \lambda_i = 0$
Multiple Rank F-test	$H_0 : \sum_{j=1}^q \delta_j \neq 0$	$H_0 : \sum_{j=1}^q \lambda_j \neq 0$	$H_0 : \sum_{j=1}^q \delta_j \neq 0$ $H_0 : \sum_{j=1}^q \lambda_j \neq 0$	$H_0 : \sum_{j=1}^q \delta_j = 0$ $H_0 : \sum_{j=1}^q \lambda_j = 0$

EMPIRICAL RESULTS

According to Granger and Newbold (1974), and Phillips (1986) the regression result may be spurious if the estimated variables are non-stationary. Thus, prior to Granger causality tests, it is necessary for this study to perform the unit root test to avoid the spurious causality problem. In this study, we carried out the KPSS unit root test to determine the order of integration for each series. The KPSS test results are reported in Table 3, and revealed that the two interested variables ($\ln S_t$ and $\ln G_t$) are non-stationary at level or integration of order one, $I(1)$, i.e., both series are stationary in first difference form ($\Delta \ln S_t$ and $\Delta \ln G_t$). These results supported the Nelson and Plosser (1982) notion that most of the macroeconomic series are non-stationary at level but become stationary after first differencing. With these findings we can proceed to the causality tests with the first difference variables (except for the MWALD causality test) to examine the causal relationship between savings and economic growth in Malaysia.

Table 3: The Results of KPSS Unit Root Test

Variables	Test statistics	
	η_μ	η_τ
Level:		
$\ln S_t$	0.976*	0.148**
$\ln G_t$	0.942*	0.216*
First difference:		
$\Delta \ln S_t$	0.160	0.150**
$\Delta \ln G_t$	0.250	0.143***

Note: The asterisks *, ** and *** denote the significance at 1%, 5%, and 10% levels respectively. The η statistics refer to the KPSS test the stationarity null hypothesis against the alternative hypothesis of a unit root. The subscripts μ and τ indicate the models that allow for drift terms and both a drift and deterministic trend respectively. The following asymptotic critical values are obtained from Kwiatkowski et al. (1992).

Significance Level:	Level:	Trend:
1%	0.739	0.216
5%	0.463	0.146
10%	0.437	0.119

Thornton and Batten (1985) postulated that causality tests are very sensitive to the lag structure incorporated into the testing equations as different lag structures may produce different causality results. In most time series analyses, estimating optimal lag structure is a crucial econometric exercise. It is also well known that most of the economic series are time series in nature and follow the autoregressive process. Koutsoyiannis (1977) documented that in a time series analysis, the parameters are implausible to be estimated when there are large numbers of lags in a small sample as there will be an inadequate degree of freedom to carry out the traditional statistical tests. Furthermore, it is almost certain that multicollinearity problems will arise and this will cause the statistical test to be insignificant at the conventional critical level. According to Enders (2004), too few lags in a regression model may cause the regression residuals not to behave like white noise. On the other hand, too many lags in a regression model will reduce the forecasting performance of the fitted model. If quarterly data is used, he suggested to begin with 12 lag as a maximum lag structure. Therefore, the present study performed a series of lag structure sequentially from 12 to 1 and the Akaike's Information Criterion (AIC) was used to choose the optimal lag structure owing to its superior performance in a small sample study (Liew, 2004). The calculated F-statistics for causality tests together with the optimal lag structure are reported in Table 4. At the 10% significant level, we found that both the parametric and nonparametric causality tests consistently reject the null hypotheses of non-causal link between savings and economic growth in Malaysia over the analysis period. This indicated that savings and economic growth in Malaysia Granger causes each other (i.e. bilateral causality) regardless of the causality tests used.

Table 4: The Results of Five Causality Tests

Type of Causality Test	$S_t \rightarrow G_t$				$G_t \rightarrow S_t$				Causal Inference
	Lag	Lead	F-Statistics	Minimum AIC	Lag	Lead	F-Statistics	Minimum AIC	
Granger	5	-	2.583**	-	11	-	4.169*	-	$S_t \leftrightarrow G_t$
Modified Sims	4	4	2.469***	-	4	4	2.581**	-	$S_t \leftrightarrow G_t$
Hsiao - AIC	(5,12)	-	-	$(-4.21 > -4.31)$	(12,11)	-	-	$(-2.50 > -2.88)$	$S_t \leftrightarrow G_t$
Modified Wald	6	-	2.818**	-	12	-	3.021*	-	$S_t \leftrightarrow G_t$
Multiple Rank F-test	12	-	2.571**	-	12	-	2.099***	-	$S_t \leftrightarrow G_t$

Note: The asterisks *, ** and *** denotes the significance at 1%, 5%, and 10% levels respectively. Each of the lag and lead in the VAR system are selected by Akaike’s information criterion (AIC). We use $d_{max} = 1$ for MWALD test. The unequal lag order is used in the Hsiao’s causality test.

The consistent bilateral causality evidence from the five causality tests highlighted two remarkable points: firstly, using different causality techniques may not yield different causality inferences. Thus, our empirical evidence showed that causality technique plays no role in explaining the inconsistent causal link between savings and economic growth. This evidence contradicted Chowdury’s (1987) postulation. A plausible explanation for the contrary result is that Chowdury (1987) did not take into account the optimal lag structure accommodated into the VAR system, thus causality results may differ. In view of literature, the financial system in Malaysia has undergone a remarkable transformation in terms of range of institutions in the system of commercial banks to unit trusts, merchant banks, and discount houses. The progress in the development of the Malaysian financial structure has also reflected the effectiveness of the financial section in mobilising savings. In addition to that, the government formed the Post Office Bank and National Savings Bank to mobilise savings from the small depositors in the rural areas. Therefore, the evidence of a bilateral causal relationship between savings and economic growth in Malaysia is not a surprising phenomenon.

Secondly, since the five causality tests are consistently suggesting that savings is an engine for economic growth in Malaysia, we may surmise that the financial system in Malaysia has successfully translated savings into productive

sectors which foster economic growth. Therefore, mobilising domestic saving is helpful rather than harmful to the Malaysian economy. In addition, this result corroborates with the findings of Tang (2008) that savings and economic growth for Malaysia Granger causes each other.

CONCLUDING REMARKS

This study employed various causality tests to detect the causal relationship between savings and economic growth in Malaysia. In particular, this study attempted to investigate whether the causal inference between savings and economic growth in Malaysia is sensitive to the particular causality test employed to ascertain the causal relationship. A remarkable finding emerged from this study is that the causality test results consistently suggest a bilateral causal relationship between savings and economic growth in Malaysia over the period of 1991:Q1 to 2006:Q3. This implied that the causal relationship between savings and economic growth in Malaysia is not sensitive to the particular causality test used in testing for causality. In other words, this study found that causality technique does not affect the causality results between savings and economic growth in Malaysia, thus the inconsistency of savings-growth nexus for Malaysia is not a result of the type of causality test. Interestingly, the causality test results has consistently affirmed that savings is a prominent

source for economic growth in Malaysia, hence mobilising the domestic resources, such as savings, is of paramount importance for further economic development in Malaysia. This finding is consistent with Lewis (1955) and endogenous growth theory which postulated that higher the savings rate will increase the rate of investment, which eventually leads to economic development and growth.

END NOTES

- ¹ In this article, Southeast Asia refers to Indonesia, Malaysia, the Philippines, Singapore and Thailand.
- ² In our study, we used parametric (i.e. Granger model, Hsiao model, modified Sims model, and MWALD) and nonparametric (i.e. Multiple F-rank) tests to examine the possibility of presence causal link between savings and economic growth in Malaysia
- ³ quation (6) will be estimated if the standard Granger causality test is adopted.
- ⁴ The standard Hsiao's test uses the Final Prediction Error (FPE) to draw the conclusion, but in this study we preferred to use Akaike's Information Criterion (AIC) due to its simplicity.
- ⁵ Some of the existing studies used unit root test to identify the order of d_{\max} .

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